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Dairy Product Consumption and Risk of Non-Hodgkin Lymphoma: A Meta-Analysis

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Abstract

Many epidemiologic studies have explored the association between dairy product consumption and risk of non-Hodgkin lymphoma (NHL), but the results remain inconclusive. A meta-analysis was performed in PubMed, Web of Science and Embase for relevant articles published up to October 2016. Pooled relative risks (RR) with 95% confidence intervals (CIs) were calculated using a random-effects model. Twenty-two articles (23 studies) with 20,960 cases and 1,151,196 participants were eligible for this meta-analysis. Pooled RRs (95% CIs) of NHL for the highest vs. lowest category of the consumption of total dairy products, milk, butter, cheese, ice cream and yogurt were 1.20 (1.02, 1.42), 1.19 (0.98, 1.45), 1.11 (0.96, 1.28), 1.06 (0.91, 1.24), 1.04 (0.89, 1.21) and 0.96 (0.82, 1.12), respectively. In subgroup analyses, the positive association between total dairy product consumption and risk of NHL was found among case-control studies (RR = 1.41, 95% CI: 1.17–1.70) but not among cohort studies (RR = 1.04, 95% CI: 0.93–1.16). Similarly, a positive association was found between milk consumption in Asian countries (RR = 2.01, 95% CI: 1.24–3.26) but not in European countries or North America, and 1.24 (1.09, 1.40) for cohort studies in the United States in 2015 [1] in spite of the considerable public health significance, the etiology of NHL remains unclear. Only a few risk factors have been established for NHL, including age, sex, and the development of NHL [2,3,4]. To date, several epidemiologic studies have examined the association between dairy product consumption and the risk of NHL. However, the results are inconsistent [5,6,7,8,9,10,11,12,13,14,15,16,17,18,19,20,21,22,23,24,25,26,27,28]. Some studies found that consumption of dairy products was associated with a decreased risk of NHL [5,6,7,8,9,10,11,12,13,14,15,16,17,18,19,20,21,22,23,24,25,26,27,28], while others found that dairy product consumption was associated with an increased risk of NHL [5,6,7,8,9,10,11,12,13,14,15,16,17,18,19,20,21,22,23,24,25,26,27,28]. Further, some studies found that dairy product consumption was not associated with the risk of NHL [5,6,7,8,9,10,11,12,13,14,15,16,17,18,19,20,21,22,23,24,25,26,27,28]. Therefore, it is necessary to further investigate the associations between specific types of dairy product consumption and the risk of NHL. Moreover, the results are inconsistent between specific types of dairy product consumption and the risk of NHL, which may be due to the small sample size of individual studies. Therefore, we conducted a meta-analysis to investigate the association of total dairy product and milk and the risk of NHL, respectively.

Keywords: dairy product; milk; cheese; yogurt; butter; ice cream; meta-analysis

1. Introduction

Dairy product is an important part of diet in many countries around the world. Nowadays many essential nutrients, such as fats, proteins, minerals, vitamin D and other bioactive nutrients. Dairy products can increase the content of total body and lumbar spine bone mineral in children [29] and decrease the risk of osteoporosis [30] and fracture [31] in the elderly. Consumption of dairy products can also reduce the risk of stroke [32] and colorectal cancer [33]. However, consumption of dairy product may be associated with several adverse health effects, for instance Parkinson's disease [34] and prostate cancer [35]. Therefore, it is necessary to further explore the effects of dairy product on non-Hodgkin lymphoma (NHL).

NHL refers to a heterogeneous group of lymphoid neoplasms that arise from B cells, T cells, or natural killer cells [36]. It is the most common hematological malignancy in the United States and the sixth most common cancer worldwide [37,38]. In the United States in 2015 [1], in spite of the considerable public health significance, the etiology of NHL remains unclear. Only a few risk factors have been established for NHL, including age, sex, and the development of NHL [2,3,4]. To date, several epidemiologic studies have examined the association between dairy product consumption and the risk of NHL. However, the results are inconsistent [5,6,7,8,9,10,11,12,13,14,15,16,17,18,19,20,21,22,23,24,25,26,27,28]. Some studies found that consumption of dairy products was associated with a decreased risk of NHL [5,6,7,8,9,10,11,12,13,14,15,16,17,18,19,20,21,22,23,24,25,26,27,28], while others found that dairy product consumption was associated with an increased risk of NHL [5,6,7,8,9,10,11,12,13,14,15,16,17,18,19,20,21,22,23,24,25,26,27,28]. Further, some studies found that dairy product consumption was not associated with the risk of NHL [5,6,7,8,9,10,11,12,13,14,15,16,17,18,19,20,21,22,23,24,25,26,27,28]. Therefore, it is necessary to further investigate the associations between specific types of dairy product consumption and the risk of NHL. Moreover, the results are inconsistent between specific types of dairy product consumption and the risk of NHL, which may be due to the small sample size of individual studies. Therefore, we conducted a meta-analysis to investigate the association of total dairy product and milk and the risk of NHL, respectively.

2. Materials and Methods

2.1. Literature Search Strategy

We conducted a literature search to identify relevant available articles published in English from PubMed, Web of Science and Embase up to October 2015. Search terms included "dairy" or "milk" or "butter" or "cheese" or "ice cream" or "yogurt" and "non-Hodgkin lymphoma" or "NHL" or "lymphoma" or "lymphoid neoplasm" or "tumor". We also reviewed the reference lists of the included studies for additional relevant studies.

2.2. Inclusion Criteria

The inclusion criteria are as follows: (1) case-control or cohort study published as an original article; (2) exposure of interest was dairy product consumption; (3) the outcome was NHL; (4) the study provided estimates of relative risk (RR) with 95% confidence interval (CI) or data to calculate these; and (5) the most recent and complete study was selected if data from the same population had been published more than once.

Two investigators (Jia Wang and Xuteng Li) searched and reviewed all identified studies independently. If the two investigators disagreed about the eligibility of an article, it was resolved by consensus with a third reviewer (Dongfang Zhang).

2.3. Data Extraction

The following information was extracted from each study by two investigators independently: first author's name, publication year, country in which the study was conducted, study design, follow-up duration, age range or mean age at baseline, sample size and number of cases, dietary assessment method, the type of dairy product, the category of dairy product consumption, and RR with 95% CI for the highest vs. lowest category of dairy product consumption. The most fully adjusted RR was selected as the summary variable adjusted for as many confounders as possible.

For dose-response analysis, the number of cases and participants (person-years) and RR (95% CI) for each category of total dairy product and milk were extracted. The median or mean level of total dairy product and milk intake for each category was used to estimate the RR for a given level of intake. If the boundaries of the highest category was not provided, we supposed the highest category width was the same as adjacent one. If median were reported as quintiles [e.g., Q1 (Q2 + Q3 + Q4 + Q5)/4], we estimated the mean for each category by averaging the median of the two adjacent quintiles. If median were reported as deciles [e.g., Q1 (Q2 + Q3 + Q4 + Q5 + Q6 + Q7 + Q8 + Q9 + Q10)/9], we estimated the mean for each category by averaging the median of the two adjacent quintiles. If median were reported as tertiles [e.g., Q1 (Q2 + Q3)/2], we estimated the mean for each category by averaging the median of the two adjacent tertiles. If the mean was not reported in the original publication, we estimated the mean for each category by using the midpoint of the upper and lower boundaries. If the upper boundary of the highest category was open, we used the lower boundary as the approximate value. If the lower boundary of the lowest category was open, we used the upper boundary as the approximate value. If studies did not report enough information to calculate the RR and 95% CI, we contacted the corresponding author to request the original data. If the original data were not available, we used the most adjusted RR and 95% CI to calculate the mean and variance of the natural log-transformed RR [39]. Two investigators (Jia Wang and Xuteng Li) extracted and reviewed all identified studies independently. If the two investigators disagreed about the eligibility of an article, it was resolved by consensus with a third reviewer (Dongfang Zhang).

2.4. Statistical Analysis

Pooled measure was calculated as the inverse variance weighted mean of the logarithm of RR with 95% CI to stabilize the variances and to minimize the influence of studies with small sample sizes [40]. Heterogeneity among RRs was tested with Cochran's Q statistic [41]. The Bonferroni-Holm method was used to correct the p-value of the Q statistic [42]. The I^2 was adopted to assess the proportion of total variation among studies [43]. A value of I^2 > 50% was considered to indicate substantial heterogeneity. Accordingly, RRs were estimated with maximum likelihood estimates of random-effects models [44]. We performed subgroup analyses to explore the potentially important covariates that might exert substantial impacts on between-study heterogeneity. Sensitivity analyses were conducted to assess the influence of individual studies on the pooled estimates by omitting studies [45]. Subgroup analyses were performed by study design (cohort studies and the studies were conducted and dietary assessment method. Influence analysis was performed with one study removed at a time to assess the influence of individual data on the pooled RR [46]. Begg's funnel plot [47] and Egger's test [48] were used to examine the publication bias.

For dose-response analysis, a two-stage, random-effects, dose-response meta-analysis [49] was performed to compute the trend from the correlated log RR estimates across levels of dairy product and milk consumption, respectively. In the first stage, a restricted cubic spline model with knots at the 10th, 50th, and 90th percentiles of the distribution [50] was fitted, taking into account the correlation within each set of published log RR estimates, using the vce(correlation) option in the xbrcspline within each set of published RR [51]. Then the study-specific estimates were combined using the restricted maximum likelihood method in a multivariate random-effects meta-analysis [52]. We assessed the nonlinearity of the relationship by using the null hypothesis that the coefficients of the second spline is equal to 0.

All statistical analyses were performed with STATA version 12.0 (Stata Corporation, College Station, TX, United States). All reported probabilities (p-values) were two-sided with a p < 0.05 considered statistically significant.

3. Results

3.1. Literature Search and Study Characteristics

We identified 1598 articles by literature search, 177 of which were excluded after review of titles and abstracts (Figure 1). One additional article was found through the reference lists of included articles. Two articles with duplicate data from the same population, one article on the association between dairy product consumption and NHL risk [53] and one article on the association between milk and NHL risk [54], were excluded because the most recent and complete data of the total dairy product consumption and the risk of NHL, and seven articles did not report RR and 95% CI [55,56,57,58,59,60,61]. 30 published articles [5,7,8,9,10,11,12,13,14,15,16,17,18,19,20,21,22,23,24,25,26,27,28,53,54,55,56,57,58,59,60,61] were eligible for this meta-analysis.

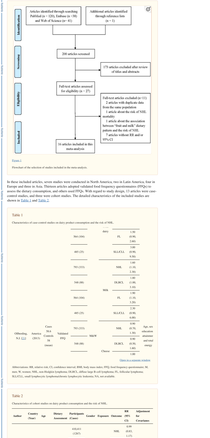

Figure 1

Flowchart of the selection of studies included in the meta-analysis.

In these included articles, seven studies were conducted in North America, two in Latin America, five in Europe and three in Asia. Thirteen articles adopted validated food frequency questionnaires (FFQs) to assess dietary intake, and nine articles used a single FFQ. Twenty-two articles (23 studies) were from cohort studies, and three were cohort studies. The detailed characteristics of the included studies are shown in Table 1 and Table 2.

Table 1

Characteristics of case-control studies on dairy product consumption and the risk of NHL.

Author	Country (Year)	Source	Dietary Assessment	Participants (Cases)	Gender	Exposure	Matched RR (95% CI)	RR Adjustment for Other Covariates
Olfberg, N.I. [12]	Sweden (2015)	Cases from the Swedish Lymphoma Register and controls from the Swedish Population Register	Validated FFQ	1016 (508)	Both sexes	Milk	RR (95% CI)	Age, sex, residential area, and energy intake
						Yes (≥3 servings/day)	1.00 (ref)	
						No (<3 servings/day)	1.00 (0.80, 1.25)	
						Fermented milk		
						Yes (≥3 servings/day)	0.83 (0.66, 1.04)	
						No (<3 servings/day)	1.00 (ref)	
						Cheese	1.00 (0.80, 1.25)	
						Yes (≥3 servings/day)	0.83 (0.66, 1.04)	
						No (<3 servings/day)	1.00 (ref)	
						Butter		
						Yes (≥3 servings/day)	1.00 (0.80, 1.25)	
						No (<3 servings/day)	1.00 (ref)	
						Ice cream		
						Yes (≥3 servings/day)	1.00 (0.80, 1.25)	
						No (<3 servings/day)	1.00 (ref)	
						Yogurt		
						Yes (≥3 servings/day)	1.00 (0.80, 1.25)	
						No (<3 servings/day)	1.00 (ref)	

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Abbreviations: RR, relative risk; CI, confidence interval; NHL, non-Hodgkin lymphoma; FFQ, Food Frequency questionnaire; NS, not Hodgkin lymphoma; DLBCL, diffuse large B-cell lymphoma; FL, follicular lymphoma; SLL/CLL, small lymphocytic lymphoma/chronic lymphocytic leukemia; TCL, peripheral T-cell lymphoma; MZBCL, marginal zone B-cell lymphoma; MM, multiple myeloma; MALT, mucosa-associated lymphoid tissue; NOS, not otherwise specified; LymphoSIGN, Lymphoma Study Association (LYSA)-led French Innovative Leukemia Organization (FILO) Group Network for Innovation Against Malignant Lymphoma; SIR, standardized incidence ratio; PY, person-years; NA, not available.

Table 2

Characteristics of cohort studies on dairy product consumption and the risk of NHL.

Author	Country (Year)	Source	Dietary Assessment	Participants (Cases)	Gender	Exposure	Matched RR (95% CI)	RR Adjustment for Other Covariates
Ji, J. [5]	Sweden (2012)	Swedish Twin Registry	Validated FFQ	61,433 (1595)	Both sexes	Milk		Age, sex, energy intake, BMI, smoking status, and family history of cancer
						Yes (≥3 servings/day)	1.00 (ref)	
						No (<3 servings/day)	1.00 (0.80, 1.25)	
						Fermented milk		
						Yes (≥3 servings/day)	0.83 (0.66, 1.04)	
						No (<3 servings/day)	1.00 (ref)	
						Cheese	1.00 (0.80, 1.25)	
						Yes (≥3 servings/day)	0.83 (0.66, 1.04)	
						No (<3 servings/day)	1.00 (ref)	
						Butter		
						Yes (≥3 servings/day)	1.00 (0.80, 1.25)	
						No (<3 servings/day)	1.00 (ref)	
						Ice cream		
						Yes (≥3 servings/day)	1.00 (0.80, 1.25)	
						No (<3 servings/day)	1.00 (ref)	
						Yogurt		
						Yes (≥3 servings/day)	1.00 (0.80, 1.25)	
						No (<3 servings/day)	1.00 (ref)	

View in a separate window

Abbreviations: RR, relative risk; CI, confidence interval; NHL, non-Hodgkin lymphoma; FFQ, Food Frequency questionnaire; NS, not Hodgkin lymphoma; DLBCL, diffuse large B-cell lymphoma; FL, follicular lymphoma; SLL/CLL, small lymphocytic lymphoma/chronic lymphocytic leukemia; NA, not available.

3.2. Quantitative Synthesis

The main results are summarized in Table 3.

Table 3

Summary risk estimates for the associated between dairy product consumption and the risk of NHL and NHL subtypes.

Exposure	Outcome	Subgroup	No. of Studies	No. of Cases	Pooled RR (95% CI)	$P_{heterogeneity}$	I^2 (%)
Total dairy product	All studies		23	20,960	1.20 (1.02, 1.42)	<0.1	0.001
					Study design		
		Cohort studies	20	20,586	1.04 (0.93, 1.16)	0.5	0.000
		Case-control studies	3	374	1.41 (1.17, 1.70)	0.5	0.000
					Study location		
		North America	8	344	1.15 (0.88, 1.49)	0.2	0.000
		Europe	11	20,197	1.06 (0.98, 1.15)	NA	NA
		Asia	3	81	0.66 (0.38, 1.15)	NA	NA
		Latin America	1	338	1.24 (0.89, 1.73)	NA	NA
		NS studies	1	1,000	1.13 (0.86, 1.49)	NA	NA
		NA studies	1	1,000	1.13 (0.86, 1.49)	NA	NA
		All studies	3	374	1.41 (0.17, 1.70)	0.5	0.000
		Specific type of dairy product					
		Case-control	3	374	1.41 (1.17, 1.70)	0.5	0.000
		Cohort	20	20,586	1.04 (0.93, 1.16)	0.5	0.000
		Specific type of dairy product					
		Cohort	20	20,586	1.04 (0.93, 1.16)	0.5	0.000
		Case-control	3	374	1.41 (1.17, 1.70)	0.5	0.000

Figure 6

The funnel plot of milk consumption and the risk of NHL. Each dot represents a different study.

4. Discussion

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This meta-analysis assessed associations between the consumption of total dairy product and specific types of dairy product and the risk of NHL, respectively. Findings from this meta-analysis showed positive associations between the consumption of total dairy product, milk, butter and ice cream and the risk of NHL, respectively. In subgroup analyses that were stratified by study design, only in case-control studies was the consumption of total dairy product and milk associated with an increased risk of NHL. In subgroup analyses that were stratified by the continent in which the studies were conducted, milk consumption was associated with an increased risk of NHL among studies conducted in North America, and cheese consumption was associated with an increased risk of NHL among studies conducted in Europe. When studies were stratified by the dietary assessment method, the consumption of milk and cheese was associated with an increased risk of NHL in those studies that used validated FFQs but not for other dietary assessment methods. Dose-response analysis suggested NHL risk increased by 5% (1.05 (1.00–1.10)) and 6% (1.06 (0.99–1.13)) for each 200 g/day increment of total dairy product and milk consumption, respectively.

The mechanisms underlying the association between dairy product consumption and the risk of NHL have been postulated from several aspects. Dairy product is rich in protein and fat. Some experimental evidence from animal studies suggested that excessive intake of protein could give rise to chronic hyperstimulation of the immune system [59]. Several animal studies have shown that the changes of animal fat and protein in diet can cause impaired immune function [60], which is one of the few well-established risk factors for NHL [61]. An animal study with rats found an increased risk of lymphomas after augmentation of the diet with casein, the major protein of milk [62]. In addition, some epidemiologic studies [63,64] suggested that polychlorinated dibenzo-p-dioxins (PCDDs) and polychlorinated dibenzofurans (PCDFs) could increase the risk of NHL. Several studies indicated that dairy was one of the most important contributors to total PCDDs and PCDFs intake in different regions [65,66,67]. Dairy product is also the main source of calcium. Several studies reported that high calcium intake could restrict the bioavailability of vitamin D [68,69]. An experimental study showed that 1,25-dihydroxyvitamin D_3, the active form of vitamin D, had an anti-proliferative and pro-differentiation effect on cells that possess vitamin D receptors in follicular lymphoma cell lines [70]. However, a meta-analysis indicated that higher vitamin D did not play a protective role in the risk of NHL [71]. So far, the effect of vitamin D on the association between dairy product consumption and the risk of NHL is still uncertain. More research is needed to explore this effect. In this meta-analysis, an inverse but not significant association was found between yogurt consumption and the risk of NHL. The anticancer effect of yogurt may depend on the large amounts of lactic acid bacteria, which can enhance the immune response of the host and exert antioxidative and antiproliferative activity [72,73].

Between-study heterogeneity is common in meta-analysis [74]. It is necessary to explore the potential sources of between-study heterogeneity. Our meta-analysis showed different levels of between-study heterogeneities in the analyses of total dairy product and specific types of dairy product consumption and the risk of NHL. In meta-regression, we found that study design and the dietary assessment method were the contributors to between-study heterogeneities of the analysis of total dairy product and cheese consumption, respectively. Case-control studies are more susceptible to recall and selection biases than cohort studies. Recall bias is the most common and inevitable bias in all case-control studies, and it refers to the differential recollection of dairy product consumption between cases and controls. The most common selection bias is that the controls are from the hospital. These potential biases may contribute to between-study heterogeneity. Subgroup analysis by dietary assessment method showed that cheese consumption can increase the risk of NHL in studies that adopted validated FFQs. The validated FFQs can provide more accurate dietary intake and make the results more credible. However, meta-regression did not find the sources of between-study heterogeneities in other analyses. The factors accounting for the heterogeneity between studies are complicated. First, the methodologies among the included studies were different. In addition, the consumption of dairy products was measured in different units across the studies included in this meta-analysis. For total dairy product consumption, the lowest intake categories ranged from <3 servings/week to 2.4 servings/day, and the highest intake categories ranged from ≥1 servings/day to 8.5 servings/day. For milk consumption, the lowest intake categories ranged from 0 to <0.5 servings/week, and the highest intake categories ranged from ≥4 servings/week to ≥14 servings/week. Third, NHL refers to a heterogeneous group of lymphomas with different prognoses and possible different etiology. The proportion of different NHL subtypes in the included studies may be different. Fourth, the consumption amount and type of dairy product largely differed between different countries and Eastern countries. All of these factors may contribute to the between-study heterogeneity in concert.

To our knowledge, this is the first meta-analysis to explore the associations between the consumption of total dairy product and specific types of dairy product and the risk of NHL. A major strength of this meta-analysis is the large number of cases included, increasing the statistical power of the study to detect the associations. Second, RRs that reflected the greatest degree of control for potential confounders were extracted, indicating that the results were more credible. Third, considering the potential differences of components in different types of dairy product, we further assessed the effects of specific types of dairy product on the risk of NHL. Fourth, dose-response analysis was conducted to explore the relationships between the consumption of total dairy product and milk and the risk of NHL, respectively. Fifth, almost all the included studies used a validated FFQ, which ensured the credibility of dietary assessment.

However, there are some limitations in this meta-analysis. First, our results mainly came from case-control studies. There were only three cohort studies included in this meta-analysis, and the positive association between total dairy product consumption and the risk of NHL was not statistically significant in cohort studies. In addition, two other cohort studies [75,76] without RR and/or 95% CI were not included in our meta-analysis. One [55] of them indicated that the consumption of more than two glasses of milk per day could increase the risk of NHL, but another [76] found no statistically significant association between the consumption of a cup of dairy product and the risk of NHL. Therefore, more cohort studies with complete data are needed to further verify these results. Second, although major confounders had been adjusted for in most of the included studies, unmeasured and residual confounding may exist. Confounders were adjusted for in each study were also different, which might affect the observed association. Third, the numbers of studies on the consumption of butter, yogurt and ice cream were limited. Thus, these cases reduced the statistical power to detect a statistically significant association. Fourth, because of the limited number of studies, we could not evaluate the association between the consumption of dairy product with different fat content and the risk of NHL. Fifth, because of the limited number of studies which examined the associations with specific histopathological subtypes of NHL, we only explored the associations between dairy product consumption and the risk of DLBCL, follicular lymphoma (FL), and small lymphocytic lymphoma/chronic lymphocytic leukemia (SLL/CLL). These studies and cases reduced the statistical power to detect statistically significant associations. Sixth, in dose-response analysis, we assumed that the amplitude of the implied exposure is same as the contiguous category. This may be imprecise given the asymmetric distributions of food consumption. Considering that most of the studies in dose-response analysis were conducted in the US, we adopted standard units reported in the United States Department of Agriculture National Nutrient Database for Standard Reference to convert our data for all the included studies. These may affect the observed association to some extent.

5. Conclusions

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In summary, this meta-analysis suggested that dairy product, but not yogurt, may increase the risk of NHL. The risk of NHL increased by 5% and 6% for each 200 g/day increment of total dairy product and milk consumption, respectively. The results mainly came from case-control studies, and these cohort studies focusing on specific types of dairy product consumption are needed to confirm the conclusion.

Acknowledgments

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The conduct of this study was not funded.

Author Contributions

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Jia Wang and Dongfeng Zhang designed the study. Jia Wang and Xutong Li carried out the literature search and data extraction. Jia Wang and Xutong Li carried out the statistical analysis. Jia Wang and Dongfeng Zhang were involved in drafting and revising the manuscript. All authors reviewed and approved the manuscript.

Conflicts of Interest

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The authors declare no conflict of interest.

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